

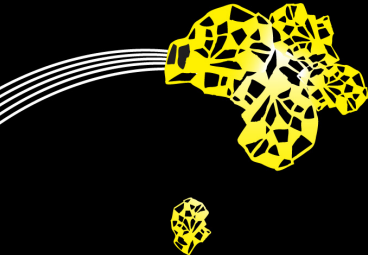
Non-iterative estimation via
method-of-moments and Croon's
method

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Overview

- 1 A non-linear factor model
- 2 Non-iterative method-of-moments
- 3 Croon's method
- 4 Monte Carlo simulation

Non-linear factor model: structural model

Example of a polynomial/non-linear factor model:

$$\begin{aligned}\eta_3 = & \gamma_1\eta_1 + \gamma_2\eta_2 + \\ & \gamma_{11}(\eta_1^2 - 1) + \gamma_{22}(\eta_2^2 - 1) + \\ & \gamma_{12}(\eta_1\eta_2 - E(\eta_1\eta_2)) + \zeta_3\end{aligned}\tag{1}$$

All latent variables are standardized, and $E(\zeta_3|\eta_1, \eta_2) = 0$.

Non-linear factor model: measurement model

Each latent variable is measured by at least two indicators and each block of indicators is connected to one LV only:

$$\mathbf{y}_i = \lambda_i \eta_i + \epsilon_i. \quad (2)$$

The indicators are standardized and the measurement errors are mutually independent and independent of the η s. The correlation of the indicators of one block can be calculated as

$$E(\mathbf{y}_i \mathbf{y}_i') = \lambda_i \lambda_i' + \Theta_i, \quad (3)$$

where the covariance matrix of the measurement errors Θ_i is a diagonal matrix.

The correlation between the indicators of different block ($i \neq j$):

$$E(\mathbf{y}_i \mathbf{y}_j') = \rho_{ij} \lambda_i \lambda_j'. \quad (4)$$

How to estimate this type model

The literature suggests several ways to estimate such models, e.g.,

- ▶ Latent Moderated Structural Equations (LMS)
[Klein, A. & Moosbrugger, H., 2000]
- ▶ Quasi-Maximum Likelihood (QML)
[Klein, A. & Muthén, B. O., 2007]
- ▶ Consistent Partial Least Squares (PLSc)
[Dijkstra, T. K. & Schermelleh-Engel, K., 2014]
- ▶ Product Indicator Approach
[Kenny, D. A. & Judd, C. M., 1984]
- ▶ Two-stage Method-of-Moments (2SMM)
[Wall, M. M. & Amemiya, Y., 2000]
- ▶ non-iterative method-of-moments
[Dijkstra, T. K., 2014, Schuberth, F. et al., in progress]
- ▶ ...

Proxies and weights

To estimate the model parameters, we build proxies for each LV as weighted linear combination of its indicators. The weights used to build proxy i are obtained as

$$\hat{\mathbf{w}}_i \propto \sum_{j \neq i} e_{ij} \mathbf{S}_{ij} \mathbf{w}_j, \quad (5)$$

where \mathbf{w}_j is an arbitrary vector of the same length as \mathbf{y}_j and $e_{ij} = \text{sign}(\mathbf{w}_j' \mathbf{S}_{ij} \mathbf{w}_j)$. Both weight vectors \mathbf{w}_j and $\hat{\mathbf{w}}_i$ are scaled: $\mathbf{w}_j' \mathbf{S}_{jj} \mathbf{w}_j$ and $\hat{\mathbf{w}}_i' \mathbf{S}_{ii} \hat{\mathbf{w}}_i = 1$.

The probability limit of $\hat{\mathbf{w}}_i$ is

$$\text{plim}(\hat{\mathbf{w}}_i) = \bar{\mathbf{w}}_i = \lambda_i / \sqrt{\lambda_i' \boldsymbol{\Sigma}_{ii} \lambda_i}. \quad (6)$$

Consistent estimation of the loadings

Calculate a factor \hat{c}_i such that the squared Euclidean difference between the off-diagonal elements of

$$\mathbf{S}_{ii} \text{ and } \hat{c}_i \hat{\mathbf{w}}_i \hat{c}_i \hat{\mathbf{w}}_i' \quad (7)$$

is minimized. As a result, we obtain

$$\hat{c}_i = \sqrt{\frac{\hat{\mathbf{w}}_i' (\mathbf{S}_{ii} - \text{diag}(\mathbf{S}_{ii})) \hat{\mathbf{w}}_i}{\hat{\mathbf{w}}_i' (\hat{\mathbf{w}}_i \hat{\mathbf{w}}_i' - \text{diag}(\hat{\mathbf{w}}_i \hat{\mathbf{w}}_i')) \hat{\mathbf{w}}_i}}. \quad (8)$$

The factor loadings can be consistently estimated as $\hat{\lambda}_i = \hat{c}_i \hat{\mathbf{w}}_i$.

The probability limit of \hat{c}_i is denoted as $\bar{c}_i = \text{plim}(\hat{c}_i) = \sqrt{\lambda_i' \Sigma_{ii} \lambda_i}$.

Relationship between latent variables and proxies

We define a population proxy:

$$\bar{\eta}_i = \bar{\mathbf{w}}_i' \mathbf{y}_i = (\bar{\mathbf{w}}_i' \boldsymbol{\lambda}_i) \eta_i + \bar{\mathbf{w}}_i' \boldsymbol{\epsilon}_i = Q_i \eta_i + \delta_i, \quad (9)$$

where Q_i (quality) is the correlation between the proxy and the latent variable. The δ_i 's have zero mean and are mutually independent and independent of the η_i 's.

Replacing $\boldsymbol{\lambda}_i$ by $\bar{c}_i \bar{\mathbf{w}}_i$, we obtain

$$Q_i = \bar{\mathbf{w}}_i' \boldsymbol{\lambda}_i = \bar{c}_i \bar{\mathbf{w}}_i' \bar{\mathbf{w}}_i. \quad (10)$$

We can estimate the quality by

$$\hat{Q}_i = \hat{c}_i \hat{\mathbf{w}}_i' \hat{\mathbf{w}}_i. \quad (11)$$

Relationship between latent variables and proxies

Relationship between the proxies' correlation and latent variables' correlation:

$$E(\bar{\eta}_i \bar{\eta}_j) = E((\bar{\mathbf{w}}_i' \mathbf{y}_i)(\bar{\mathbf{w}}_j' \mathbf{y}_j)) = \quad (12)$$

$$E((\bar{\mathbf{w}}_i' \boldsymbol{\lambda}_i \eta_i + \bar{\mathbf{w}}_i' \boldsymbol{\epsilon}_i)(Q_j \eta_j + \delta_j)) = \quad (13)$$

$$E(Q_i Q_j \eta_i \eta_j) + E(Q_i \eta_i \delta_j) + E(\delta_i Q_j \eta_j) + E(\delta_i \delta_j) = \quad (14)$$

$$Q_i Q_j E(\eta_i \eta_j), \quad (15)$$

where $E(\bar{\eta}_i \bar{\eta}_j)$ is estimated by the sample covariance between the proxies i and j .

Interaction terms

Starting point is a model with a two-way interaction term:

$$\eta_3 = \gamma_1\eta_1 + \gamma_2\eta_2 + \gamma_{12}(\eta_1\eta_2 - E(\eta_1\eta_2)) + \zeta_3. \quad (16)$$

The γ 's can be obtained by solving the following moment equations:

$$\begin{pmatrix} E(\eta_1\eta_3) \\ E(\eta_2\eta_3) \\ E(\eta_1\eta_2\eta_3) \end{pmatrix} = \begin{pmatrix} 1 & E(\eta_1\eta_2) & E(\eta_1^2\eta_2) \\ & 1 & E(\eta_1\eta_2^2) \\ & & E(\eta_1^2\eta_2^2) - E(\eta_1\eta_2)^2 \end{pmatrix} \begin{pmatrix} \gamma_1 \\ \gamma_2 \\ \gamma_{12} \end{pmatrix}, \quad (17)$$

where the moments are given by:

$$E(\bar{\eta}_i\bar{\eta}_j) = Q_i Q_j E(\eta_i\eta_j), \quad (18)$$

$$E(\bar{\eta}_i^2\bar{\eta}_j) = Q_i^2 Q_j E(\eta_i^2\eta_j), \quad (19)$$

$$E(\bar{\eta}_i^2\bar{\eta}_j^2) = Q_i^2 Q_j^2 (E(\eta_i^2\eta_j^2) - 1) + 1, \text{ and} \quad (20)$$

$$E(\bar{\eta}_i\bar{\eta}_j\bar{\eta}_k) = Q_i Q_j Q_k E(\eta_i\eta_j\eta_k). \quad (21)$$

Distributional assumptions?

So far, **no distributional assumptions are necessary**. We only assume that

- ▶ the moments exist,
- ▶ the measurement errors are mutually independent and independent from the η 's, and
- ▶ that the structural error term is independent from the η 's of the right-hand side.

Quadratic terms

Adding quadratic terms to the equation with the one-way interaction term:

$$\eta_3 = \gamma_1 \eta_1 + \gamma_2 \eta_2 + \gamma_{11}(\eta_1^2 - 1) + \gamma_{12}(\eta_1 \eta_2 - E(\eta_1 \eta_2)) + \gamma_{22}(\eta_2^2 - 1) + \zeta_3. \quad (22)$$

Now higher moments are required, and therefore more assumptions are necessary, in particular, assumptions about the higher moments of the error terms

$$E(\bar{\eta}_i^3) = Q_i^3 E(\eta_i^3) + E(\delta_i^3), \quad (23)$$

$$E(\bar{\eta}_i^4) = Q_i^4 E(\eta_i^4) + 6Q_i^2(1 - Q_i^2) + E(\delta_i^4), \text{ and} \quad (24)$$

$$E(\bar{\eta}_i^3 \bar{\eta}_j) = Q_i^3 Q_j E(\eta_i^3 \eta_j) + 3E(\bar{\eta}_i \bar{\eta}_j)(1 - Q_i^2). \quad (25)$$

Additional assumptions are required

Way out: we assume that δ_i , has the same higher moments as the normal distribution, i.e.,

$$E(\delta_i^3) = 0, \text{ and} \tag{26}$$

$$E(\delta_i^4) = 3 [\text{var}(\delta_i)]^2 = 3(1 - Q_i^2)^2. \tag{27}$$

Presentation of Yves Rosseel & Ines Devlieger

Why we may not need SEM after all

Yves Rosseel & Ines Devlieger
Department of Data Analysis
Ghent University – Belgium

March 15, 2018
Meeting of the SEM Working Group – Amsterdam

future plans and challenges

- challenge: (analytical) standard errors that perform well in the presence of missing indicators and/or non-normal (but continuous) indicators
- challenge: categorical indicators
- challenge: nonlinear/interaction effects (involving latent variables)
- challenge: models where the distinction between the measurement part and the structural part of the model is not clear
- solved: extension to multilevel SEM (see talk by Ines on EAM in Jena)
- future plans: study the relationship with other related approaches:
 - consistent PLS
 - model-implied instrumental variables estimation
 - two-step approaches
 - ...

Croon's approach

Croon's approach [Croon, M. A., 2002] can be used to estimate non-linear factor models by conducting the following steps:

- ▶ Estimate each measurement model by CFA to obtain the factor loading estimates $\hat{\lambda}_i$.
- ▶ Build proxies by sum scores, i.e., unit weights $\hat{w}_i = \mathbf{1} / \sqrt{\mathbf{1}' \mathbf{S}_{ii} \mathbf{1}}$.
- ▶ Estimate the quality of the proxy as $\hat{Q}_i = \hat{w}_i' \hat{\lambda}_i$.
- ▶ Estimate the moments using these quality estimates.

Design of the Monte Carlo simulation

- ▶ Structural model:

$$\eta_3 = 0.3\eta_1 + 0.4\eta_2 + 0.12(\eta_1^2 - 1) + 0.15(\eta_1\eta_2 - 0.3) + 0.1(\eta_2^2 - 1) + \zeta_3. \quad (28)$$

- ▶ Correlation between η_1 and η_2 is set to $\rho_{12} = 0.3$
- ▶ Each latent variable is measured by 3 indicators, $\lambda'_i = (0.9 \quad 0.85 \quad 0.8)$
- ▶ Exogenous variables are normally distributed
- ▶ Sample size of $N = 400$ and 500 runs
- ▶ Estimators: Croon's approach, Non-iterative method-of-moments, and LMS

Results of the Monte Carlo simulation

Para.	true	Croon ¹	Non-iter. ²	LMS ¹
γ_1	0.300	0.297 (0.048)	0.297 (0.048)	0.297 (0.047)
γ_2	0.400	0.404 (0.047)	0.403 (0.047)	0.402 (0.045)
γ_{11}	0.120	0.120 (0.045)	0.119 (0.044)	0.117 (0.041)
γ_{12}	0.150	0.148 (0.063)	0.146 (0.062)	0.147 (0.059)
γ_{22}	0.100	0.106 (0.043)	0.105 (0.043)	0.101 (0.039)
ρ_{12}	0.300	0.299 (0.052)	0.300 (0.052)	0.299 (0.052)

¹No inadmissible solutions were produced

²Inadmissible solutions are removed, therefore the results are based on 484 estimations

Outlook

What we have done so far in case of the non-iterative method-of-moments estimator:

- ▶ Implementation in R (cSEM package [↗](#))
- ▶ Allow for correlated measurement errors within a block of indicators
- ▶ Assume normality of all exogenous variables \Rightarrow facilitates calculation of the moments

For future research:

- ▶ Generation of non-normally distributed data maintaining the covariance structure
- ▶ Estimate non-recursive models, e.g., by 2SLS
- ▶ Deal with categorical indicators
- ▶ Apply approach to other methods, e.g., MIIV-SEM
- ▶ Test for overall model fit

Thank you!
Questions/Comments?

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